

The New Demographic Transition: Most Gains in Life Expectancy Now Realized Late in Life[†]

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The original “demographic transition” describes a process that began in Europe by the early 1800s with decreases in mortality followed, usually after a lag, by decreases in fertility (Davis 1945; for an overview in this journal, see Lee 2003). According to Lee and Reher (2011, p. 1), “this historical process ranks as one of the most important changes affecting human society in the past half millennium.” The increase in life expectancy associated with this demographic transition has been accompanied by rising levels of per capita output, which have in turn spurred further improvements in population health through better nutrition and living standards (Fogel 1994; Barker 1990) and, especially since World War II, through advances in medical care (in this journal, Cutler, Deaton, and Lleras-Muney 2006). At the same time, increases in life expectancy have resulted in a higher proportion of each cohort living long enough to participate in the production of goods and services. Reductions in fertility are also closely linked to higher labor force participation rates among women (Galor and Weil 1996; Costa 2000; Guinnane 2011).

During the original demographic transition, mortality decline prior to fertility decline often led to larger cohorts concentrated in working ages; this transitional change in the age structure of the population provided a boost to income that has

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[†]To access the Appendix, visit
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been called a “demographic dividend” (Bloom, Canning, and Sevilla 2003). Swift (2011) documents a significant two-way positive relationship between life expectancy and GDP per capita between 1820 and 2001 for 13 high-income countries.

Now, the United States and many other countries are experiencing a new kind of demographic transition. Instead of additional years of life being realized early in the lifecycle, they are now being realized late in life. At the beginning of the twentieth century, in the United States and other countries at comparable stages of development, most of the additional years of life were realized in youth and working ages; and less than 20 percent was realized after age 65. Now, more than 75 percent of the gains in life expectancy are realized after 65—and that share is approaching 100 percent asymptotically. The choice of age 65 to illustrate this new demographic transition is somewhat arbitrary, but if we used 60 or 70 instead, the results would be qualitatively similar.

The new demographic transition is a *longevity* transition: How will individuals and societies respond to mortality decline when almost all of the decline will occur late in life? This issue is broader and more far-reaching than the issue of cohort size in each age group, with its usual focus on the prospective retirement of the unusually large “baby boomer” cohort, and has important socioeconomic implications independent of patterns of fertility.

When the gains in life expectancy occur mainly towards the end of life, they contribute more to the age bracket that is traditionally mostly retired rather than to the age bracket in prime working years. Retirees are highly dependent on transfers from the working population for living expenses, including large consumption of medical care. Thus, gains in life expectancy concentrated at the end of life can unsettle an economy’s balance between production and consumption in ways that pose a long-run challenge for public policy. The obvious changes needed (at least “obvious” to many economists) would be to raise productivity, the savings rate, and the age of retirement, but how to accomplish such goals is controversial and uncertain.

This paper covers the years 1900–2007 for the United States and 16 other “developed countries,” chosen for the continuity of their mortality data: Australia, Belgium, Canada, Denmark, England and Wales, Finland, France, Iceland, Italy, Netherlands, Northern Ireland, Norway, Scotland, Spain, Sweden, and Switzerland. We focus on demographic statistics including life expectancy at birth and at age 65, the percent of each birth cohort expected to survive to age 65, and the share of the increase in life expectancy at birth realized after age 65. For the U.S. economy, we also calculate expected labor force participation for each birth cohort, which allows us to investigate how changes in mortality affect labor force participation and work-life as a share of life expectancy. Results on the longevity transition and expected labor force participation for the United States and other high-income countries are followed by consideration of economic and social changes in China and other countries that are experiencing an earlier stage of the original demographic transition. The paper concludes with a brief discussion of the long-run implications of the new demographic transition.

The Longevity Transition

To examine long-term trends in life expectancy at birth, we draw upon the life tables in the Human Mortality Database, which offers high-quality demographic data for selected countries and regions compiled by a respected group of demographers at (<http://www.mortality.org>). We first extract data on life expectancy at birth; in particular, we calculate “period” life expectancy, which is the projected average age of death for a cohort if it experienced the age-specific death rates prevailing at the year of birth. We also look at rates of survival from birth to age 65 and life expectancy at age 65. We use the five-year period life tables since 1900 (or earliest available year) for each of the 17 countries or regions in the Human Mortality Database that have data extending back at least 70 years. The five-year intervals help to smooth annual fluctuations in demographic trends.

We calculate changes for nine overlapping 20-year intervals: 1907–1927, 1917–1937, and so on up to 1987–2007.¹ (The years ending in “7” are chosen to represent mid points of each of our five-year intervals.) To calculate the change in years lived past 65, we first multiply survival to 65 by life expectancy at age 65 for each five-year period and then take differences across 20-year intervals. Finally, we calculate the change in years lived past 65 as a percentage of change in life expectancy at birth for each country for each of the nine 20-year intervals.

Figure 1A shows that life expectancy at birth has increased almost continuously for well over a century in high-income countries. Much of this rise in life expectancy was due to a particularly large fall in death rates for infants, children, and young adults, resulting in a sharp rise in the percentage of a cohort surviving to age 65, as indicated in Figure 1B. Survival rates from birth to age 65 more than doubled over the twentieth century from 40.9 percent in 1900–04 to 83.3 percent in 2005–09 in the United States. Similarly, survival rates from birth to age 65 in 16 high-income comparators increased from 42.0 to 87.8 percent over the same period.

The other major demographic change that contributes to the longevity transition is an increase in life expectancy at age 65, an increase which has become larger in recent decades as shown in Figure 2A. The interaction between the increase in life expectancy at age 65 and the increase in the percentage of the cohort that survives to age 65 has resulted in an exceptionally large increase in the share of the gain in life expectancy that is realized after age 65. As can be seen in Figure 2B, that share was only about 20 percent during each 20-year period at the beginning of the twentieth

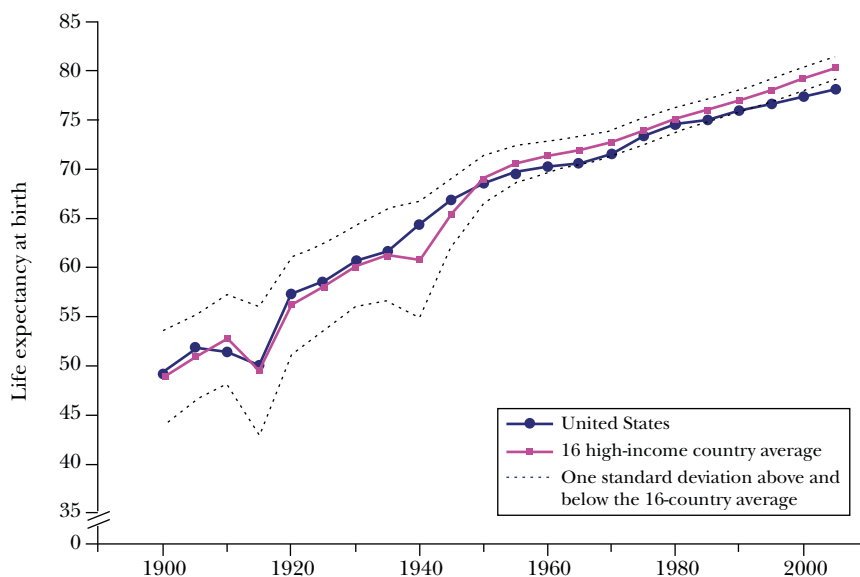
¹ For our detailed underlying data on the five-year averages for each country, see the online Appendix with this paper at (<http://e-jep.org>). Online Appendix tables 1–3 show the decreases in the coefficient of variation across the 17 high-income countries for the demographic variables portrayed in Figures 1 and 2. To include data for the United States prior to 1933 (when the Human Mortality Database series begins for the United States), we use life table data from U.S. National Vital Statistics Reports, derived from death registration states for the period 1900 to 1928, and for the whole United States thereafter (all races combined). For a small share of observations at the beginning of the century—Australia, Canada, and Northern Ireland in 1900–1919; Spain in 1900; and the United States in 1905, 1915, and 1925—we use imputed values from regressions with year and country fixed effects and country-specific linear time trends.

Figure 1

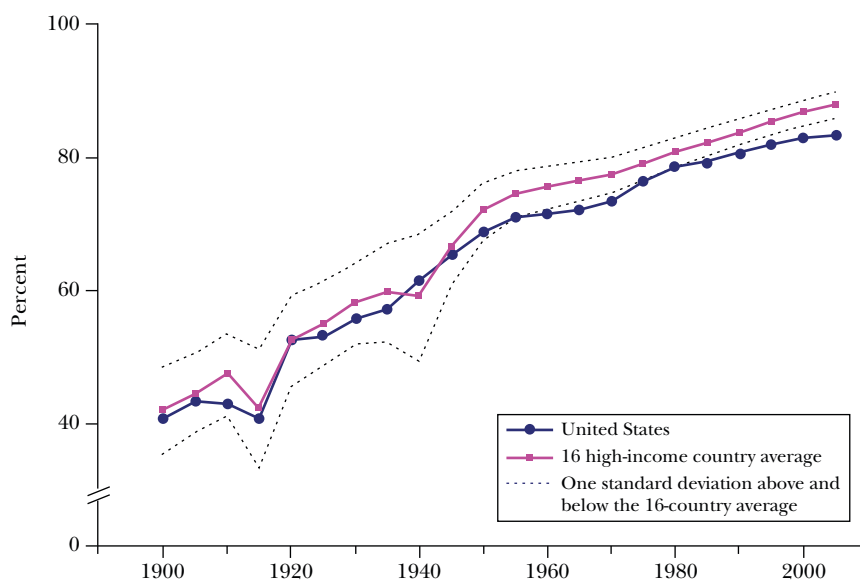
Life Expectancy at Birth and Survival to Age 65, since 1900

(in the United States and 16 other high-income countries)

A: Life Expectancy at Birth

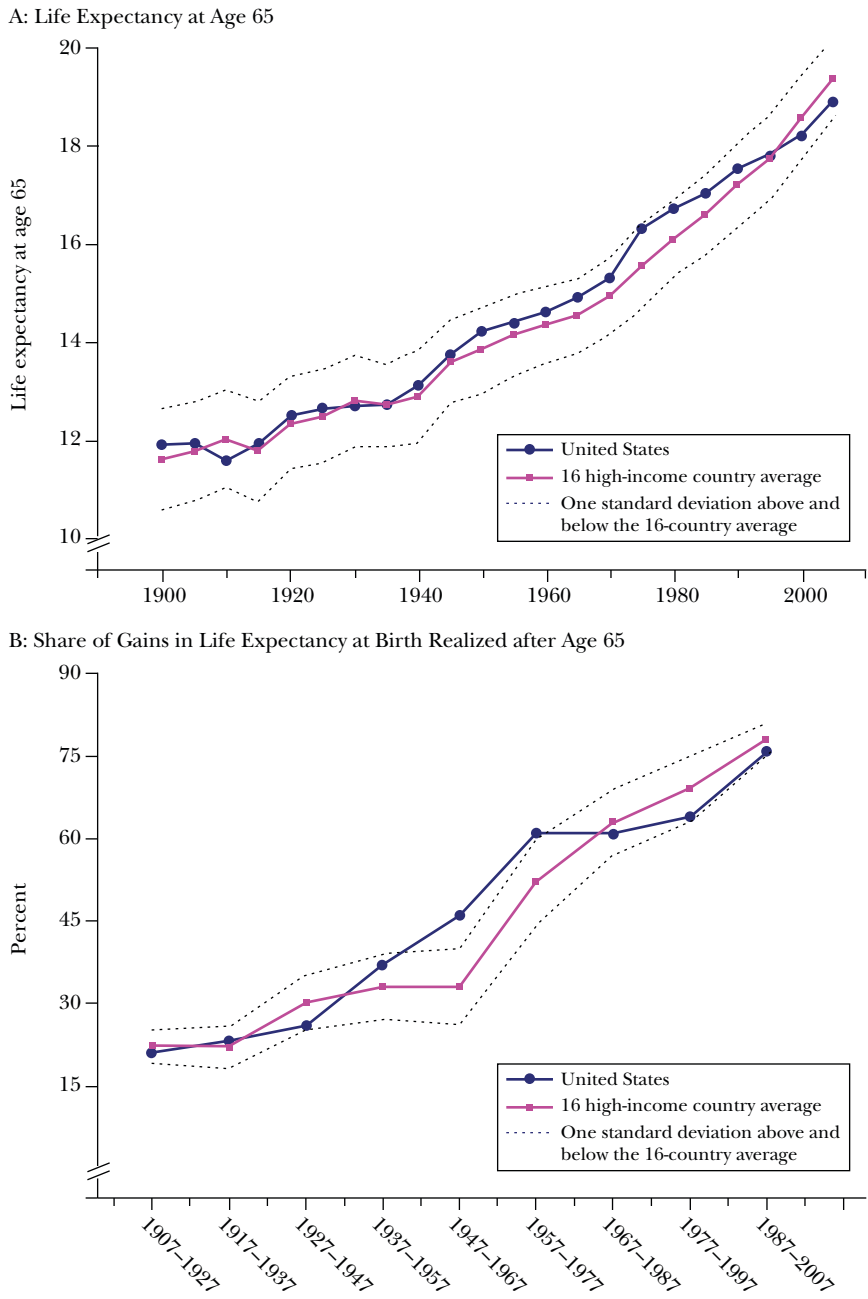


B: Percent of Birth Cohort Expected to Survive to Age 65



Source: Authors' calculations using data from the Human Mortality Database and other sources as detailed in the online Appendix.

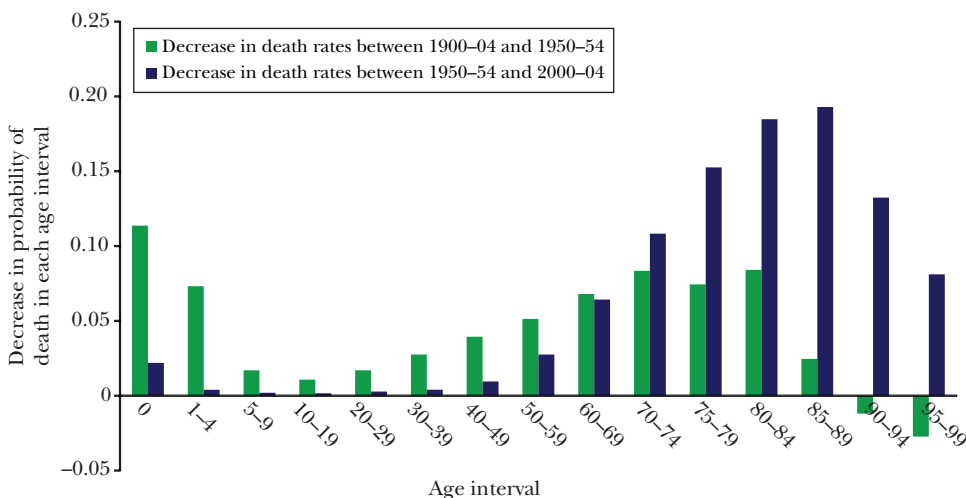
Figure 2
Life Expectancy at Age 65 and Gains in Life Expectancy Realized after Age 65 Since 1990
(in the United States and 16 other high-income countries)



Source: Authors' calculations using data from the Human Mortality Database and other sources as detailed in the online Appendix.

Figure 3

Decrease in Death Rates by Age Group in England and Wales, 1900–1904 to 1950–54 and 1950–1954 to 2000–2004



Source: Authors' calculations using data from the Human Mortality Database.

century, but it was 76 percent in the United States and 78 percent for the 16-country mean by the end of the century, and is approaching 100 percent asymptotically. Our results here are quite similar to, and extend over time, those of Lee and Tuljapurkar (1997) based on the 1995 survival profile of the United States.

We can illustrate the shift in survival improvement toward older ages by comparing the age distribution of mortality decline between the first half and second half of the twentieth century for a region with particularly reliable long-run data: England and Wales. Figure 3 shows that between 1900–1904 and 1950–1954, declines in death rates were largest for infants and children, whereas between 1950–54 and 2000–2004, declines were most salient for those over age 70. This pattern of age-specific mortality decline across the twentieth century was similar for Sweden, another country where reliable long-run data is available.²

The actual survival of a given birth cohort will differ from the estimates of life expectancy at birth when survival is changing over time. Remember, estimates of life expectancy at birth (what we earlier called “period” life expectancy) are based on the age-specific death rates prevailing at that year of birth. For example, in 1900–04, life expectancy at birth in England and Wales was 48.6 years. In contrast, the cohort born in 1900–1904 had a cohort life expectancy (actual mean age of death) of 53.8 years, since they experienced part of the increase in survival shown

² For details on Sweden, see the online Appendix. Figure 3 shows a slight increase in death rates for the oldest [90+] age groups between 1900–1904 and 1950–1954, perhaps because of small numbers, less-reliable data, and/or survival of a less-healthy cohort to those ages.

in Figures 1–3. The cohort born only 17 years later experienced a cohort life expectancy of 62.4 years, whereas “period” life expectancy at birth did not reach that level until 1935–1939.³

Nevertheless, we find that estimates based on cohort life tables prepared by the Social Security Administration (Bell and Miller 2005) exhibit a similar trend towards survival gains realized late in life: for men, the share of life expectancy increases realized after age 65 was 28 percent between the 1900 and 1920 birth cohorts, rising to a projected 62 percent between the 1980 and 2000 birth cohorts. For women, the share of life expectancy gains realized after age 65 increased from 30 percent (between the 1900 and 1920 birth cohorts) to an estimated 69 percent (between the 1980 and 2000 birth cohorts).

The century-long demographic trends shown in Figures 1 and 2 have been similar in all 17 countries with available data. From a U.S. perspective, the main difference is lagging survival to 65 compared to the other 16 countries (the U.S. line is below the 16-country average in Figure 1B); also, the United States experienced a larger rise in female life expectancy at age 65 between the 1940s and 1970s than the other countries. The relative differences among countries have decreased over time, especially for life expectancy at birth and survival to age 65.

The Longevity Transition and Expected Labor Force Participation

One of the most significant economic effects of the longevity transition is on expected lifetime labor force participation, partly in terms of total years in the workforce and especially in terms of years in the workforce as a fraction of expected years of life. Two factors affecting the connection from life expectancy to years of work are 1) whether the growing numbers of elderly are healthy enough to work and 2) the economic, social, and political pressures for a period of retirement at the end of life.

Greater longevity can have opposing effects on age-specific health status. If improved survival is correlated with reductions in morbidity for the elderly, then illness may be compressed into the end of life, as posited by the “compression of morbidity” hypothesis (Fries 1980). On the other side, medical interventions do tend to keep alive those who are in worse health (Zeckhauser, Sato, and Rizzo 1985), which suggests the possibility that the longer-lived elderly could be sicker for a longer period. The net effect of rising longevity on age-specific morbidity is an empirical question. According to the National Long-Term Care Survey, the share of elderly Americans with severe disabilities decreased from 26.2 to 19.7 percent between 1982 and 1999 (Manton and Gu 2001). Milligan and Wise (2011) find a

³ Survival gains have been so dramatic that period and cohort survival significantly differs. For example, age-specific death rates for England and Wales in 1900–1904 would have led to only 43.7 percent of women and 36.4 percent of men surviving to 65. But of the cohort born in 1900–1904, 61.3 percent of women and 49.6 percent of men actually survived to age 65.

strong within-country correlation between declining mortality and improved self-assessed health for several European countries. Thus, the empirical record suggests that better health in terms of both improved survival and reduced morbidity could tend to raise age-specific rates of labor force participation. Changes in occupational structure which lower the physical demands of work also can increase participation.

Higher incomes tend to increase the demand for leisure, in the form of fewer hours of work per week and, especially recently, as a block at the end of life (Costa 1998; Murphy and Topel 2006). Furthermore, several factors might give rise to a negative interaction between improved survival and employment, at least for some subgroups. For example, the reduced selection effect of mortality might also increase the proportion of the cohort that is less valued in employment (because of less stamina, ambition, education, and the like), reducing age-specific labor force participation. Alternatively, if firms have pyramid-like organizational structures with many jobs at entry and fewer at higher levels in the hierarchy—such as the military’s “up or out” policy regarding age and promotion of officers—then increases in survival will lead to crowding at higher levels of the pyramid and lower rates of participation. Moreover, a sharp rise in employment rates for women, at wages that were often below those paid to men, might have led to some decrease in the demand for men’s labor.

On net, which of these forces have predominated over the past century, and which are likely to predominate in the future? Estimates of what we call “expected labor force participation” can help answer this question.

Calculating Expected Labor Force Participation

We define “expected labor force participation” (XLFP) as the total years an individual is expected to participate in the labor force, based on period estimates of survival, and labor force participation by gender and age. That is

$$XLFP_{jt} = \sum_{i=1}^{100} \pi_{ijt} L_{ijt},$$

where L_{ijt} is the labor force participation rate for age i and gender j in year t , weighted by probability of survival to age i (π_{ijt}). It is necessary to examine men and women separately because of the large upsurge in female labor force participation between the 1950s and 2000 (Goldin 1986, 1990; Costa 2000). Our calculations rely on labor force participation rates from decennial censuses (1900–1930) and the Current Population Survey (1942–2007). As in the earlier estimates of life expectancy, we can calculate both “period” expected labor force participation, which is based on the age-specific labor force participation rates prevailing at a certain point of time, or the actual realized labor force participation rates for a birth cohort; these estimates will differ when age-specific labor force participation rates are changing over time.

Changes in lifetime expected labor force participation can be decomposed into two factors: changes in survival to given ages and changes in age/sex-specific

rates of labor force participation. For example, we calculate the effect of improving survival, holding age-specific labor force participation rates constant at their 2007 values. We also calculate the effect of changing rates of labor force participation, holding survival rates constant.⁴

Our work is related to the literature on expected lifetime work hours (Hazan 2009) and work-life expectancy (Smith 1982), including the work-life estimates for the U.S. population from the 1950s through the early 1980s from the Bureau of Labor Statistics.⁵ As far as we are aware, this paper is the first to produce work-life estimates for the United States covering the period 1900 to 2007, decompose those changes into survival and age/sex-specific labor force participation effects, and to estimate work-life expectancy relative to life expectancy at birth for a broader range of countries in recent decades.

U.S. Expected Labor Force Participation since 1900

In the early twentieth century, most of the increase in life expectancy arose from the dramatic decrease in mortality at young ages. This change first increased the years of youth dependency for these cohorts, and then increased expected labor force participation—the expected number of years that an individual will be in the labor force if he or she participates at the average labor force participation rate for each sex and age in a given year.

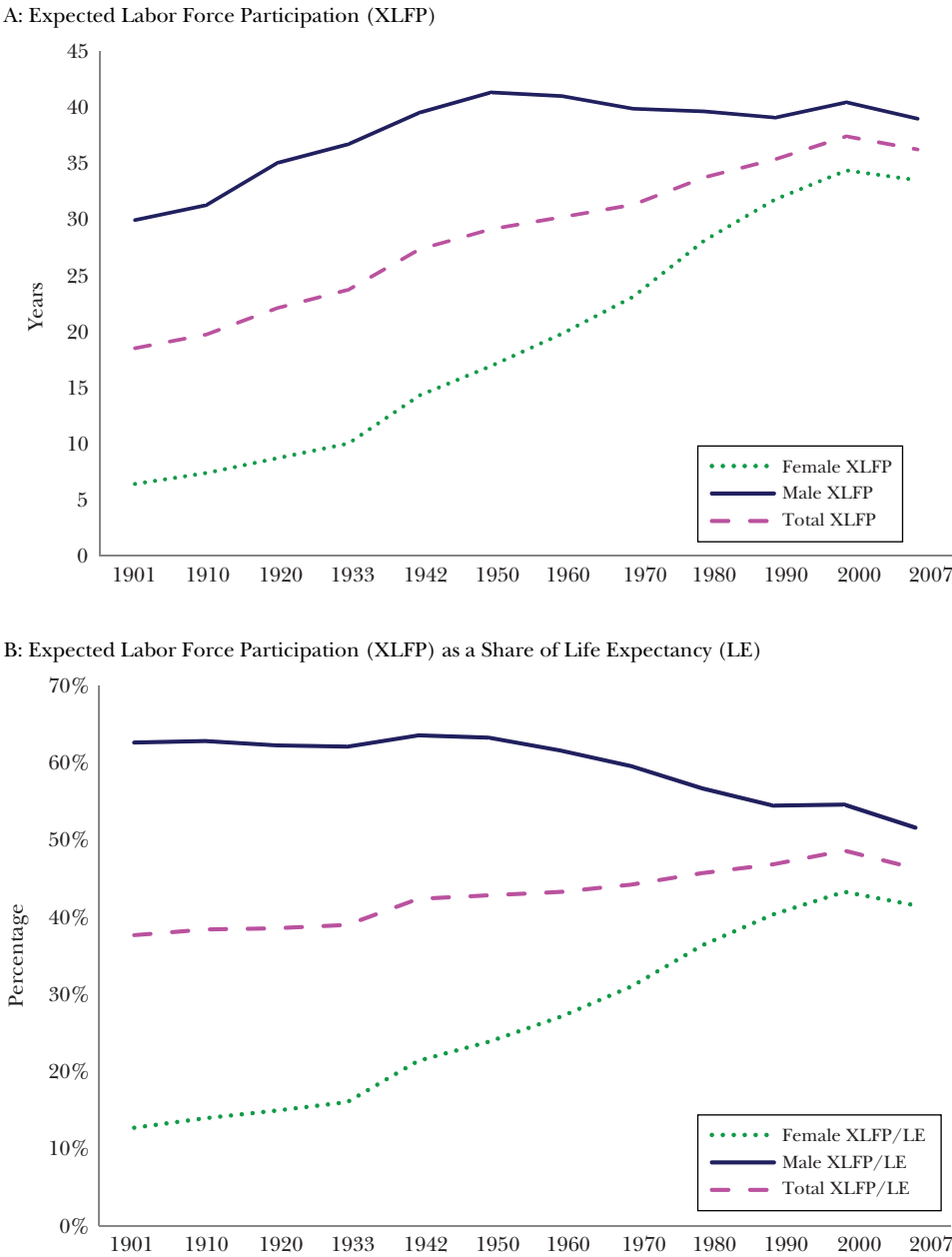
Figure 4A shows that years of expected labor force participation at birth for U.S. males increased by a third—from about 30 to 40 years—between 1900 and 1950. For the most recent half century, however, increases in survival have been offset by decreasing age-specific labor force participation rates for men, causing expected lifetime labor force participation to be relatively constant at about 40 years. Because life expectancy at birth has continued to increase, male expected labor force participation as a fraction of expected years of life has declined, as shown in Figure 4B. Table 1 shows that in the United States between 1900 and 2000, male labor force participation increased from 30 to 40.5 years, female participation from 6.4 years to 34.4 years, and for the total population from 18.5 to 37.4 years. This increase in years of expected labor participation is two-thirds of the total gain in life expectancy at birth of 28.2 years over the twentieth century.

How much of this change is attributable just to longer life expectancies? If we hold age-specific rates of labor force participation constant but allow survival rates to grow at the actually observed pace, the rise in life expectancy alone would have increased expected labor force participation by 13.3 years for males and by 10.8 years for females since 1900 (as shown in Table 1). The effect of mortality decline was concentrated in the first half of the twentieth century. Indeed, for men,

⁴ These are decompositions 1B and 2B, respectively, in Table 7 of the online Appendix. Alternative calculations, using 1900 as the base year (decompositions 1A and 2A), show similar results.

⁵ In other pre-existing work in this area, Hunt, Pickersgill, and Rutemiller (2001) update worklife estimates for the U.S. based on 1998–1999 labor force participation rates. Millimet, Nieswiadomy, Ryu, and Slottje (2003) use a regression framework. In related research, Hazan (2009) estimates lifetime working hours for U.S. men born between 1840 and 1970 and for the U.S. population born between 1890 and 1970.

Figure 4
U.S. Expected Labor Force Participation since 1990 and as a Share of Life Expectancy at Birth



Source: Authors' calculations using data from the Human Mortality Database and other sources as detailed in the online Appendix.

Table 1

Expected Labor Force Participation in the United States, by Sex, 1900–2007

Year	Men					Women		
	Male XLFP	Male XLFP holding LFP constant	Male XLFP adjusted for hours worked	Male XLFP/LE ₀	Male XLFP adjusted for hours/LE ₀	Female XLFP	Female XLFP holding LFP constant	Female XLFP/LE ₀
	XLFP					XLFP		
1900	30.0	25.7	37.28	62.6%	77.9%	6.4	22.7	12.7%
1910	31.3	27.1	39.96	62.8%	80.2%	7.4	24.1	13.9%
1920	35.1	30.4	37.65	62.2%	66.8%	8.7	26.3	14.9%
1933	36.7	32.3	40.40	62.0%	68.2%	10.0	28.3	16.0%
1942	39.5	34.1	42.66	63.5%	68.5%	14.3	30.1	21.3%
1950	41.3	35.6	38.22	63.2%	58.4%	16.9	31.3	23.8%
1960	41.0	36.3	36.79	61.5%	55.2%	19.8	32.0	27.0%
1970	39.9	36.4	34.67	59.5%	51.7%	23.1	32.2	31.0%
1980	39.6	37.4	n.a.	56.6%	n.a.	28.1	32.8	36.3%
1990	39.1	37.9	n.a.	54.4%	n.a.	31.8	33.1	40.3%
2000	40.5	38.7	n.a.	54.5%	n.a.	34.4	33.3	43.2%
2007	39.0	39.0	n.a.	51.6%	n.a.	33.5	33.5	41.5%
Change, 1900 to most recent	9.0	13.3	-2.6	-11.0%	-26.1%	27.1	10.8	28.8%

Year	Total (men and women)				
	Total XLFP	Total XLFP holding LFP constant	Total XLFP adjusted for hours worked	Total XLFP/LE ₀	Total XLFP adjusted for hours/LE ₀
1900	18.5	24.2	n.a.	37.6%	n.a.
1910	19.8	25.6	n.a.	38.4%	n.a.
1920	22.1	28.4	n.a.	38.5%	n.a.
1933	23.7	30.3	29.0	39.0%	47.5%
1942	27.4	32.2	29.2	42.3%	45.1%
1950	29.1	33.6	29.0	42.8%	42.5%
1960	30.2	34.2	28.8	43.2%	41.2%
1970	31.3	34.4	28.9	44.2%	40.7%
1980	33.8	35.2	n.a.	45.7%	n.a.
1990	35.4	35.6	n.a.	46.8%	n.a.
2000	37.4	36.0	n.a.	48.6%	n.a.
2007	36.3	36.3	n.a.	46.3%	n.a.
Change, 1900 to most recent	17.7	12.0	n.a.	8.7%	n.a.

Sources: Author calculations based on survival data from the Human Mortality Database (1933–2007), supplemented by data for death registration states for 1900–1920; and labor force participation rates from decennial censuses (1900–1930) and the Current Population Survey (1942–2007). Adjustments for hours worked draw from Hazan (2009). See the online Appendix for details.

Notes: Expected Labor Force Participation (XLFP) is calculated as the total years an individual is expected to participate in the labor force based on period estimates of labor force participation and survival by gender and age. XLFP for a given year represents the expected number of years that an individual would be in the labor force if he or she participates at the average LFP rate for each age in that given year. LE₀ is life expectancy at birth. “XLFP holding LFP constant” uses 2007 age- and sex-specific labor force participation rates, but allows survival to each age to vary as it actually did between 1900 and 2007.

if we hold age-specific labor force participation rates constant but allow survival rates to vary in calculating expected labor force participation (“male XLFP holding LFP constant”), the ratio of years of expected labor force participation to life expectancy at birth was relatively constant at 54 percent from early in the twentieth century until about 1970 (not shown in the table). At that point, it began a slow but seemingly inexorable decline, now falling to about 50 percent.

Actual years of expected labor force participation, reflecting both survival effects and changes in age-specific labor force participation rates, have also begun to decline. As shown in both Table 1 and Figure 4B, the ratio of years of expected labor force participation to life expectancy at birth ($XLFP/LE_0$) has declined for U.S. men from 62.6 percent in 1900 to 51.6 percent in 2007. That same ratio for women increased from 12.7 percent in 1900 to 43.2 percent in 2000, before declining slightly to 41.5 percent by 2007. For the overall U.S. population, years of expected labor force participation divided by life expectancy at birth peaked at 48.6 percent in 2000 and declined slightly to 46.3 percent by 2007.

Since 1950, increases in survival and declines in age-specific participation rates of men tended to offset one another. For example, between 1950 and 2007, labor force participation rates of men ages 45–54 declined from 95.8 percent to 88.2 percent, but survival to age 50 increased from 84.1 to 92.2 percent, so the total expected years in the labor force between ages 45 and 55 remained eight years.⁶ For women, increases in years of expected labor force participation mostly reflect increases in age-specific rates of labor force participation, especially after 1950. Accordingly, for women, if we hold age-specific labor force participation rates constant but allow survival rates to vary in calculating expected labor force participation (“female XLFP holding LFP constant”), the ratio of years of expected labor force participation to life expectancy at birth has declined slowly but steadily from about 45 percent in the first few decades of the twentieth century to about 40 percent (not shown in the table).

The increase in female labor force participation since the late 1950s could be considered primarily a one-time substitution from unpaid home production to paid work outside the home (Goldin 1990; Costa 2000). If so, then the decrease in years of expected labor force participation for women in the United States since 2000 would reflect completion of the one-off change and the beginning of a similar trend as seen for men—that is, a decline of years in the labor force as a share of life expectancy at birth.

Taking into account the decrease in the intensive margin—annual hours worked per full-time worker—tends to reinforce the conclusion that expected work life has declined as a fraction of life expectancy at birth. Hazan (2009)

⁶ For the detailed data behind these calculations across the range of ages, for both men and women, see online Appendix Table 7, which offers alternative decompositions of changes in both male and female labor force participation. Online Appendix Table 7 also shows that holding age-specific labor force participation rates constant (at either their 1900 or 2007 values) would have led to a larger increase in male expected labor force participation than actually observed.

estimated lifetime work hours over the past century conditional on survival to age five. We adapt Hazan's data to life expectancy at birth to calculate years of expected labor force participation adjusted for hours worked and show the results in Table 1 (the online Appendix available with this paper at <http://e-jep.org>) has details of our calculations).

Calculation of a century-long trend in expected years of labor force participation in other high-income countries is not possible because there is no reliable source for internationally comparable labor force participation rates before 1980. Given the similarities in trends of both survival and labor force participation across these countries for the available years, we suspect the trend of declining expected labor force participation as a share of life expectancy at birth that we found for the United States reflects a broad and robust trend that countries experience as they reach high life expectancy levels. Indeed, with the sole exception of the Netherlands, the ratio of years of expected labor force participation to life expectancy at birth has declined since 1980 for males in all other high-income countries in our analyses.⁷ Adjusting for a decline in work hours would reinforce this trend.

Demographic Transition across Stages of Economic Development

The demographic transition traces out a pathway, with many societies arrayed along earlier phases of the transition roughly and imperfectly in accordance with their per capita incomes. Many developing countries are currently experiencing the original demographic transition. For example, Table 2 shows that between 1990 and 2010, the share of years lived past 65 as a percentage of increase in life expectancy at birth was only a little over one-third in Vietnam and Brazil, and less than one-quarter in Bangladesh—comparable to levels a century earlier in today's high-income countries.

Improving health and increasing life expectancy at birth clearly can contribute to better living standards for the world's poor (World Health Organization 2002). Data on labor force participation for developing countries is not always reliably comparable across countries and over time. Nevertheless, the importance of improved survival for gains in expected labor force participation at early stages of the longevity transition can be illustrated with extant data. For example, in 1980 only 70 percent of Indonesian men survived to age 45; by 2007, 90 percent did. This improved survival added 10 years to expected labor force participation rates for Indonesian males between 1980 and 2007. As a result, expected labor force participation rates for Indonesian males rose to 43.7 years, which was 64.5 percent of life expectancy at birth in 2007.

⁷ The online Appendix tables provide calculations of expected labor force participation across 15 countries since 1980; see Appendix Table 8 in the online Appendix available with this paper at <http://e-jep.org>. Milligen and Wise (2011, p. 17) examine the age at which male mortality was 1.5 percent in 1977 and 2007, finding that at that age almost 90 percent of UK men were employed in 1977, but by 2007, only 30 percent were.

Table 2
**The Longevity Transition in Asia and Select
 Developing Countries**

<i>Country</i>	<i>Change in years lived past 65 as a percentage of change in life expectancy at birth, 1990–2010</i>	
	<i>Males</i>	<i>Females</i>
Japan	72.7%	87.0%
South Korea	45.4%	57.1%
China	51.9%	40.6%
Philippines	26.2%	36.0%
Indonesia	26.1%	35.7%
Brazil	34.2%	35.0%
Vietnam	32.5%	34.7%
India	23.6%	25.8%
Bangladesh	20.7%	25.4%

Source: Authors' calculations based on the life tables for each country prepared by the International Programs Center of the U.S. Bureau of the Census in its International Data Base.

China and India are especially important cases to consider, given their large populations and relatively rapid economic development. In India, the share of years lived past 65 as a percentage of increase in life expectancy at birth was only one-quarter (as shown in Table 2) in the most recent 20-year period. For China, that share was 52 percent for men and 41 percent for women in the 1990–2010 period.

China's position reflects the rapidity of its demographic transition since the early 1970s and its achievement of relatively high levels of health despite low per capita income by the end of the Mao era (Banister 1987; Wang 2011). Indeed, despite the higher death rates associated with the Great Leap Famine of 1959–1961, China's growth in life expectancy from approximately 35–40 years in 1949 to 65.5 years in 1980 ranks as the most rapid sustained increase in documented global history.⁸ These earlier health improvements and growth of the working-age population contributed to China's unprecedented economic growth for the past quarter-century. Wang and Mason (2008) estimate that between 1982 and 2000, about 15 percent of China's rapid growth in output per capita stemmed from the demographic dividend. (Bloom and Williamson (1998) estimate that one-quarter to one-third of the growth rates in the "East Asian miracle" stemmed from the demographic dividend.) Although the pace of mortality decline in China has slowed, it

⁸ Miller, Eggleston, and Zhang (2012) assess the relative importance of various explanations proposed for these gains, including better nutrition, widespread public health interventions, improved access to medical care, and increases in educational levels. They find that gains in education and public health campaigns jointly explain 25–32 percent of the crude death rate decline under Mao, and similar proportions of the dramatic reductions in infant and under-five mortality in that period.

continues: Chinese life expectancy increased between 1990 and 2010 from 69.9 to 76.8 years for women and from 66.9 to 72.5 years for men.

With a rapid demographic transition to relatively low mortality and low fertility, China's population is now aging (Peng 2011). Many policy challenges loom as China establishes social and economic institutions commensurate with its transition to a middle-income, market-based economy with a large elderly population (Eggleston and Tuljapurkar 2010; Chen, Eggleston, and Li forthcoming). One additional challenge for China in reducing the growth-slowing potential of the new demographic transition is China's increasing burden of chronic disease. Fueled by rapid urbanization, increases in high-fat and calorie-rich diets, reductions in physical activity, unabated male smoking and other factors, prevalence of chronic disease in China has quickly caught up with that of high-income countries. For example, the age-standardized prevalence of diabetes among adults in China was 9.7 percent in 2007–2008, more than three times reported prevalence in 1994 (Yang et al. 2010), comparable to the U.S. rate of 8.3 percent overall in 2010 and 11.3 percent among adults (CDC 2011), and higher than the OECD average (OECD 2011).

The timing and the rapidity of the longevity transition has varied across countries and regions. For example, in Japan between 1950 and 1970, only 13.1 percent of increase in male life expectancy at birth was realized after age 65; for women, that figure was 17.3 percent. During the 1990 to 2009 period, Japan led the world in the new demographic transition, with the share of gains in life expectancy at birth realized after age 65 reaching 72.7 percent for men and 87 percent for women (again, as shown in Table 2).

The original and the new demographic transitions are inextricably intertwined with the evolution of social and economic institutions (Aoki 2011). Evidence is mounting that no society at an advanced stage of economic development can presume that further gains in longevity will contribute to growth of per capita income under currently prevailing institutions. For example, Lee and Mason (2011) compare the “average age of consumption” to the “average age of labor income”⁹ across a large group of countries for which they and their international collaborators have collected detailed generational accounts, including the value of assets and transfer wealth from social support programs (but not including bequests or value of nonmarket labor). They find that for developing countries, net transfers flow strongly downward from older to younger ages. However, in a “sea change” analogous to what we call the new demographic transition, “the direction of intergenerational transfers in the population has shifted from downward to upward, at least in a few leading rich nations” including Germany, Austria, and Japan (Lee and Mason 2011, p. 116). Although the Lee–Mason estimates are cross-sectional, the link to the longevity transition is clear: for the 13 countries that overlap between

⁹ They construct the average ages of consumption and labor income as follows: “The average age of consumption is calculated by multiplying each age by the aggregate consumption at that age, summing these products over all ages, then dividing by the total amount of consumption at all ages. An equivalent calculation gives the average age of labor income” (Lee and Mason 2011, p. 123).

their dataset and ours, there is a strong negative correlation (-0.89) between the share of gains in life expectancy over the past 20 years that were realized after age 65 and the current number of years by which the average age of income exceeds the average age of consumption. In other words, the more the gains in life expectancy are concentrated in traditional retirement years, the closer the intergenerational transfers are to being upward rather than downward.

For a broader group of 107 countries, Bloom, Canning, and Fink (2010) calculate counterfactual annual growth rates of per capita income between 1960 and 2005, using 2005–2050 projections of demographics. The results vary depending on the level of economic development. They find that in most non-OECD countries, declining youth dependency would more than offset increasing old-age dependency. However, about half of countries would have grown more slowly using 2005–2050 projections of demographics. Among 26 OECD countries analyzed, 25 of them (Turkey is the exception) would have had lower economic growth—averaging 2.1 rather than 2.8 percent per year—under the counterfactual of 2005–2050 demographic change.

Policy Implications of the New Demographic Transition

Historically, adults produced more than they consumed and supported children. With such a pattern in place, the increase in proportion of the population in older years implied by the demographic transition might have been thought to shift out the social budget constraint as people expanded their number of years worked. However, “a funny thing happened along the way: societies invented retirement . . . and the economic consequences of population aging are now viewed with alarm” (Lee and Mason 2011, p. 115).

Retirement, a relatively new phenomenon in human history, can be viewed as a response to many economic and social changes. Contributing factors include the shift from self-employment on farms or small businesses to wage and salary status; more rapid technological change, resulting in more rapid obsolescence of human capital (alongside compensation packages that often underpay at the beginning and overpay at the end of a career, as discussed in Lazear 1981); the introduction of a variety of health and welfare programs which assist the elderly but also discourage work; an income-driven increase in the demand for leisure, with the diminishing marginal value of an even shorter work week overtaken by the efficiency gains of a block of leisure at the end of life; and, in times of high unemployment, public concern about job opportunities for younger workers.

Will the new demographic transition inevitably lead to slower economic growth? As people foresee longer lives, they might choose to work longer, save more, and/or invest in human capital in sufficient amounts and innovative enough ways that longer lives continue to contribute to increased prosperity. In this spirit, Bloom, Canning, and Fink (2010) assert that “the problem of population ageing is more a function of rigid and outmoded policies and institutions than a problem of demographic change per se” (p. 607).

It is not clear, however, that the United States or other high-income countries even further along in the new demographic transition are reshaping their policies and institutions sufficiently in response to the longevity transition. Although both the United States and France have increased the age of retirement or age to qualify for early retirement, social welfare systems across the high-income countries of the world continue to give strong incentives for earlier, rather than later, retirement (Gruber and Wise 1998). Between 1965 and 2005, the correlation between change in male life expectancy at birth and change in retirement age is actually negative: -0.21 (Bloom, Canning, and Fink 2010, p. 591). This trend cannot continue indefinitely: longer and longer retirement lives are not consistent with continued increases in per capita income unless there are significant increases in savings, investment, and productivity. It is ironic that the same phenomenon that led to higher GDP per capita—namely higher life expectancy—could now lead to lower GDP per capita.

Successful navigation of the new demographic transition calls for a combination of policies to give incentives for more savings and investment (including in human capital) earlier in the lifecycle and for additional work later in the lifecycle. Two forces in particular might move the society in that direction: improvement in health, and reductions in the transfers that the elderly can expect to receive from the young.

Public policy should encourage higher labor force participation for the elderly, both by reducing the disadvantages that employers face when employing older workers and by providing enhanced incentives to individuals to continue to work. “People cannot expect to finance 20–25 year retirements with 35-year careers,” Shoven noted (as quoted in Haven 2011). “It just won’t work. Not in Greece [or] the United States . . . Eventually, we are going to have to increase retirement ages.” However, increasing labor force participation for the 65-plus age group alone probably won’t make a big difference: even a doubling of those rates from their 2007 levels of 12.6 for women and 20.5 for men would not bring the U.S. ratio of expected labor force participation to life expectancy at birth back to its 2000 level. Increased labor force participation by men in the 50–64 age bracket is also needed.

Public policy might also seek to improve productivity, with an emphasis on education and building human capital early in the lifecycle, and on investment to reduce morbidity and improve the ability to work later in life. Whether compression of morbidity later in life will continue depends on whether improvements in medical technology and in the socioeconomic determinants of health are offset by adverse trends such as increasing obesity. A potentially promising focus here would be to consider investments in public health and medical technologies that reduce morbidity and improve quality of life, as well as more focus on medical innovations that reduce costs of care. (One example of a policy consistent with both objectives would be expansion of palliative care as a substitute for what can otherwise be extremely expensive end-of-life care in a hospital—especially in countries where the concept of hospice services is relatively new, such as China.)

Finally, increased savings, investment, and capital formation could help in fueling endogenous growth (Lucas 1988; Romer 1990). U.S. personal savings rates

have been low for many decades. Increasing the savings rate of individuals before they retire would ameliorate the potential adverse impact of longevity on economic growth. Countries will need to make fiscally realistic structural changes to entitlement programs—such as Medicare and Social Security in the United States—to support acceptable living standards and improvements in health.

High-income societies are now facing a new demographic transition: the longevity transition. They must decide how to respond to mortality decline when almost all of the decline will occur late in life. Additional increases in life expectancy will result in further declines in expected labor force participation as a percentage of life expectancy at birth unless there is a significant rise in labor force participation rates across both middle and older ages. Of course, increased life expectancy has great value independent of its relationship to per capita income (Murphy and Topel 2006). The original demographic transition gave society a “demographic gift” of higher per capita incomes (Bloom and Williamson 1998) without much need for a policy response, but the new demographic transition requires politically difficult policies if societies wish to preserve a positive relationship running from increased longevity to greater prosperity.

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